

Probabilities and mathematical needs

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Pascal Bootcamp 2010, 07/06/10

Outline

1 Linear Algebra

- Vector spaces
- Orthogonality, dot product, norm
- Matrices
- Determinant
- Matrix decompositions (SVD, Choleski, LU, QR)

2 Probabilities

- Vocabulary, usual laws (discrete, continuous)
- Conditional probabilities
- Bayes rule, maximum likelihood, maximum a posteriori
- Entropy, Kullback-Leibler divergence, perplexity
- Bounds

3 Optimization

- Minima, maxima, saddle points
- Convex functions
- Primal and dual problems, Lagrange multipliers

Vector spaces

Example (\mathbb{R}^n)

$$\mathbb{R}^n = \{x = (x_1, \dots, x_n)^T : x_i \in \mathbb{R} \forall i\}$$

- ▶ $x, y \in \mathbb{R}^n \Rightarrow x + y = (x_1 + y_1, \dots, x_n + y_n)^T \in \mathbb{R}^n$
- ▶ $x \in \mathbb{R}^n, \lambda \in \mathbb{R} \Rightarrow \lambda x = (\lambda x_1, \dots, \lambda x_n)^T \in \mathbb{R}^n$
- ▶ $\mathbb{R}^n = \{x : \exists (\lambda_1, \dots, \lambda_n) \in \mathbb{R}^n \text{ s.t. } x = \lambda_1 e_1 + \dots + \lambda_n e_n\}$
where $e_i = (0, \dots, 0, \underset{i}{1}, 0, \dots, 0)$.

Example (Solutions of homogeneous differential equations)

$$\mathcal{S} = \{f : \mathbb{R} \rightarrow \mathbb{R} : \forall t, f''(t) + f(t) = 0\}$$

- ▶ $f \in \mathcal{S} \Rightarrow -f \in \mathcal{S}$
- ▶ $f, g \in \mathcal{S} \Rightarrow f + g \in \mathcal{S}$
- ▶ $f \in \mathcal{S}, \lambda \in \mathbb{R} \Rightarrow \lambda f \in \mathcal{S}$
- ▶ $\mathcal{S} = \{f : \mathbb{R} \rightarrow \mathbb{R} : \exists (\lambda_1, \lambda_2) \in \mathbb{R}^2 \text{ s.t. } f = \lambda_1 \cos + \lambda_2 \sin\}$

Vector spaces

Example ($L^2(\mathbb{R})$)

$$L^2(\mathbb{R}) = \left\{ f : \mathbb{R} \rightarrow \mathbb{R} : \int_{\mathbb{R}} |f(x)|^2 dx < \infty \right\}$$

- ▶ $f \in L^2(\mathbb{R}) \Rightarrow -f \in L^2(\mathbb{R})$
- ▶ $f, g \in L^2(\mathbb{R}) \Rightarrow f + g \in L^2(\mathbb{R})$
- ▶ $f \in L^2(\mathbb{R}), \lambda \in \mathbb{R} \Rightarrow \lambda f \in L^2(\mathbb{R})$
- ▶ $L^2(\mathbb{R})$ is not the span of any finite number of its elements.
- ▶ Dot product : $f, g \in L^2(\mathbb{R}), \langle f, g \rangle = \int_{\mathbb{R}} f(x)g(x)dx$
- ▶ Norm : $\|f\|_{L^2(\mathbb{R})} = \left(\int_{\mathbb{R}} |f(x)|^2 dx \right)^{\frac{1}{2}}$
- ▶ Closeness :
 $\forall n, f_n \in L^2(\mathbb{R})$ and $\|f_n - f\|_{L^2(\mathbb{R})} \xrightarrow{n \rightarrow \infty} 0$ implies $f \in L^2(\mathbb{R})$.

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Vector spaces

Definition (Vector space)

A set \mathcal{S} is called a real vector space if it is endowed with

- ▶ an “addition” which is :
 - ▶ stable : $x, y \in \mathcal{S} \Rightarrow x + y \in \mathcal{S}$,
 - ▶ commutative and associative,
 - ▶ with an nul element $0 \in \mathcal{S}$ s.t. $\forall x \in \mathcal{S}, 0 + x = x$,
 - ▶ for which all elements are invertible $x \in \mathcal{S} \Rightarrow -x \in \mathcal{S}$.
- ▶ the multiplication by a scalar in \mathbb{R} which is :
 - ▶ stable : $x \in \mathcal{S}, \lambda \in \mathbb{R} \Rightarrow \lambda x \in \mathcal{S}$.
 - ▶ associative and distributive over '+’.

Vector spaces may be decomposed into subspaces :

Definition (Subspace)

A subset F of a vector space \mathcal{S} is called a subspace of \mathcal{S} if the previous properties are preserved in F .

Vector subspaces, family of vectors, dimension

- ▶ **Supplementary subspaces** :
 - ▶ F, G subspaces, $F \cap G = \{0\}$, $S = F + G$.
 - ▶ Any $x \in S$ has a unique decomposition $x = x_F + x_G$.
- ▶ Subspaces may be generated from a **family of vectors** :
 - ▶ $y \in \text{Span}\{x_1, \dots, x_n\}$ iff $\exists \lambda_1 \dots \lambda_n \in \mathbb{R}$ s.t. $y = \sum_{i=1}^n \lambda_i x_i$.
 - ▶ The family $\{x_i\}_{i=1..n}$ is linearly independent iff the decomposition $y = \sum_{i=1}^n \lambda_i x_i$ is unique.
 - ▶ Conversely if $F = \text{Span}\{\{x_i\}_{i=1..n}\}$ then the family $\{x_i\}_{i=1..n}$ is said to generate F .
- ▶ The **dimension** of a (sub)space F is the cardinal of its largest linearly independent family.
 - ▶ *Ex* : $\dim(\mathbb{R}^d) = d$, $\dim(\mathcal{S}_{\text{diff. eq.}}) = 2$, $\dim(L^2(\mathbb{R})) = +\infty$.
 - ▶ A **hyperplane** is a subspace of which the supplementaries have dimension 1.
 - ▶ If $\dim(S) = n$, an hyperplane is any subspace of dimension $n-1$. *Ex* : *lines in \mathbb{R}^2 , planes in \mathbb{R}^3 .*

Bases

- ▶ The family $\{x_i\}_{i=1..n}$ is a **basis** of \mathcal{S} iff it generates \mathcal{S} and is linearly independent. **Here n may be ∞ !**
 - ▶ The cardinal of any basis is exactly the dimension of \mathcal{S} (finite or not).
 - ▶ For $y \in \mathcal{S}$ there is a unique decomposition $y = \sum_{i=1..n} \lambda_i x_i$.

Example

- ▶ In \mathbb{R}^d :
 - ▶ $\{e_i\}_{i=1..d}$, where $e_i = (0, \dots, 0, \overset{i}{\uparrow} 1, 0, \dots, 0)$ is a basis.
 - ▶ $y = (y_1, \dots, y_d)^T = \sum_{i=1..d} y_i e_i$.
- ▶ In $L^2([0, 2\pi])$:
 - ▶ $\{\cos(mt), \sin(mt)\}_{m \in \mathbb{N}}$ is a basis.
 - ▶ $f \in L^2([0, 2\pi])$, $f(t) = \sum_{m \in \mathbb{N}} (a_m \cos(mt) + b_m \sin(mt))$.

Orthogonality, dot product, norm

In \mathbb{R}^d :

- ▶ The dot product is defined as :

$$\langle x, y \rangle_{\mathbb{R}^d} = \sum_{i=1}^d x_i y_i$$

- ▶ It is linked to the Euclidian norm :

$$\|x\| = \sqrt{\langle x, x \rangle_{\mathbb{R}^d}} = \sqrt{\sum_{i=1}^d |x_i|^2}$$

$$\langle x, y \rangle_{\mathbb{R}^d} = \|x\| \|y\| \cos(\theta)$$

- ▶ Any subspace has a unique orthogonal supplementary

Orthogonality, dot product, norm

Definition (norm, dot product, Hilbert space)

\mathcal{S} a vector space.

- ▶ $\|\cdot\| : \mathcal{S} \rightarrow \mathbb{R}^+$ is a **norm** iff
 1. $\|x\| = 0 \Leftrightarrow x = 0$
 2. $\lambda \in \mathbb{R}, x \in \mathcal{S}, \|\lambda x\| = |\lambda| \|x\|$
 3. $x, y \in \mathcal{S}, \|x + y\| \leq \|x\| + \|y\|$
- ▶ a **dot product** is a bilinear symmetric application of \mathcal{S}^2 to \mathbb{R} .
 - ▶ then $x \rightarrow \sqrt{\langle x, x \rangle}$ is a norm.
 - ▶ x and y are **orthogonal** when $\langle x, y \rangle = 0$.
 - ▶ F has a unique orthogonal supplementary F^\perp .
 - ▶ For any x , the unique decomposition $x = x_F + x_{F^\perp}$ also verifies : $\|x\|^2 = \|x_F\|^2 + \|x_{F^\perp}\|^2$.
- ▶ a **Hilbert space** \mathcal{H} is a vector space endowed with a dot product $\langle \cdot, \cdot \rangle_{\mathcal{H}}$, that is closed for the induced norm.

Orthonormal bases

- ▶ A basis $\{e_i\}_{i=1..n}$ is **orthonormal** of \mathcal{H} iff $\langle e_i, e_j \rangle_{\mathcal{H}} = \delta_{\{i=j\}}$.
 - ▶ $y \in \mathcal{H}$, the unique decomposition $y = \sum_{i=1..n} \lambda_i x_i$ verifies :
 1. $\lambda_i = \langle y, e_i \rangle_{\mathcal{H}}$
 2. $\|y\|_{\mathcal{H}}^2 = \sum_i |\lambda_i|^2$

Example

- ▶ In \mathbb{R}^d :
 - ▶ $\{e_i = (0, \dots, 0, \overset{i}{1}, 0, \dots, 0)\}_{i=1..d}$ is a an orthonormal basis.
 - ▶ $y = (y_1, \dots, y_d)^T = \sum_{i=1..d} y_i e_i$ and $\|y\| = \sqrt{\sum_{i=1..d} y_i^2}$.
- ▶ In $L^2([0, 2\pi])$:
 - ▶ $\{\cos(mt), \sin(mt)\}_{m \in \mathbb{N}}$ is an orthonormal basis.
 - ▶ $f \in L^2([0, 2\pi])$, $f(t) = \sum_{m \in \mathbb{N}} (a_m \cos(mt) + b_m \sin(mt))$
where $a_m = \int_0^{2\pi} f(t) \cos(mt) dt$, $b_m = \int_0^{2\pi} f(t) \sin(mt) dt$.
 - ▶ $\|f\|_{L^2}^2 = \int_0^{2\pi} |f(t)|^2 dt = \sum_{m \in \mathbb{N}} (|a_m|^2 + |b_m|^2)$.

Hyperplanes

H a hyperplane then $\dim F^\perp = 1$ hence there is a vector $u \in \mathcal{H}$ such that :

$$F^\perp = \text{Span} \{u\} = \mathbb{R}u \quad \text{and} \quad \|u\|_{\mathcal{H}} = 1.$$

- ▶ Equation of H : $H = \{x \in \mathcal{H} : \langle x, u \rangle_{\mathcal{H}} = 0\}$.

$$H = \{x = (x_1, x_2)^T : x_1 u_1 + x_2 u_2 = 0\}$$

- ▶ The distance from x to H is : $d(x, H) = |\langle x, u \rangle_{\mathcal{H}}|$.

$$d(x, H) = |x_1 u_1 + x_2 u_2|$$

- ▶ The projection of x on H is : $P_H(x) = x - \langle x, u \rangle_{\mathcal{H}} u$.

$$P_H(x) = x - (x_1 u_1 + x_2 u_2)u$$

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Matrices

- ▶ Let $H_1 = \mathbb{R}u_1^\perp$, $H_2 = \mathbb{R}u_2^\perp$, \dots , $H_m = \mathbb{R}u_m^\perp$ be m hyperplanes of \mathbb{R}^d and $F = \bigcap_{i=1}^m H_i$.
- ▶ The equation of F is a system of m linear equations with d unknowns :

$$\begin{cases} u_1^1 x_1 + u_1^2 x_2 + \dots + u_1^d x_d = 0 \\ u_2^1 x_1 + u_2^2 x_2 + \dots + u_2^d x_d = 0 \\ \vdots \\ u_m^1 x_1 + u_m^2 x_2 + \dots + u_m^d x_d = 0 \end{cases}$$

which is equivalent to the matrix-vector equation :

$$Ux = 0 \Leftrightarrow \begin{pmatrix} u_1^1 & u_1^2 & \dots & u_1^d \\ u_2^1 & u_2^2 & \dots & u_2^d \\ \vdots & \vdots & \ddots & \vdots \\ u_m^1 & u_m^2 & \dots & u_m^d \end{pmatrix} \begin{pmatrix} x_1 \\ x_2 \\ \vdots \\ x_d \end{pmatrix} = \begin{pmatrix} 0 \\ 0 \\ \vdots \\ 0 \end{pmatrix}$$

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- ▶ The equation of F is a system of m linear equations with d unknowns :

$$\begin{cases} u_1^1 x_1 + u_1^2 x_2 + \dots + u_1^d x_d = b_1 \\ u_2^1 x_1 + u_2^2 x_2 + \dots + u_2^d x_d = b_2 \\ \vdots \\ u_m^1 x_1 + u_m^2 x_2 + \dots + u_m^d x_d = b_m \end{cases}$$

which is equivalent to the matrix-vector equation :

$$Ux = b \Leftrightarrow \begin{pmatrix} u_1^1 & u_1^2 & \dots & u_1^d \\ u_2^1 & u_2^2 & \dots & u_2^d \\ \vdots & \vdots & \ddots & \vdots \\ u_m^1 & u_m^2 & \dots & u_m^d \end{pmatrix} \begin{pmatrix} x_1 \\ x_2 \\ \vdots \\ x_d \end{pmatrix} = \begin{pmatrix} b_1 \\ b_2 \\ \vdots \\ b_m \end{pmatrix}$$

Matrices

- ▶ A **matrix** in $\mathbb{R}^{m \times d}$ is an array made of m row-vectors of \mathbb{R}^d or equiv. d column vectors of \mathbb{R}^m (e.g. U).
- ▶ The matrix-vector product Ux may be seen as :
 1. Using column vectors $U^j = (u_1^j, u_2^j, \dots, u_m^j)^T$:

$$Ux = \sum_{j=1}^d x_j U^j, \quad \text{where } U^j \in \mathbb{R}^m.$$

2. Using row vectors $U_i = (u_i^1, u_i^2, \dots, u_i^d)$:

$$Ux = \begin{pmatrix} \langle U_1^T, x \rangle_{\mathbb{R}^d} \\ \langle U_2^T, x \rangle_{\mathbb{R}^d} \\ \vdots \\ \langle U_m^T, x \rangle_{\mathbb{R}^d} \end{pmatrix} \in \mathbb{R}^m$$

Note : U is a representation of a linear operator : $x \in \mathbb{R}^d \rightarrow Ux \in \mathbb{R}^m$.

Matrices

► Notation :

$$A \in \mathbb{R}^{m \times d} = \begin{pmatrix} a_{1,1} & a_{1,2} & \cdots & a_{1,d} \\ a_{2,1} & a_{2,2} & \cdots & a_{2,d} \\ \vdots & \vdots & \ddots & \vdots \\ a_{m,1} & a_{m,2} & \cdots & a_{m,d} \end{pmatrix} = (a_{i,j})_{\substack{i=1 \dots m \\ j=1 \dots d}}$$

► Operations on matrices :

► $\mathbb{R}^{m \times d}$ is a real vector space with $A + B = (a_{i,j} + b_{i,j})_{\substack{i=1 \dots m \\ j=1 \dots d}}$

► Matrix product : $A \in \mathbb{R}^{m \times p}$, $B \in \mathbb{R}^{p \times d}$, then :

$$AB \in \mathbb{R}^{m \times d} \quad \text{s.t.} \quad (AB)_{i,j} = \sum_{k=1}^p a_{i,k} b_{k,j}$$

Note : $AB \neq BA$!

► Matrix transposition : $A \in \mathbb{R}^{m \times d}$, then :

$$A^T \in \mathbb{R}^{d \times m} = (a_{j,i})_{\substack{j=1 \dots d \\ i=1 \dots m}}$$

Square matrices ($m=d$)

- ▶ Matrix product is stable in $\mathbb{R}^{d \times d}$, so some are invertible !
- ▶ Remarkable matrices
 - ▶ Diagonal matrices.

$$D = \begin{pmatrix} \lambda_1 & 0 & \cdots & 0 \\ 0 & \lambda_2 & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & \lambda_d \end{pmatrix}$$

- ▶ Upper and Lower triangular matrices :

$$U = \begin{pmatrix} u_{1,1} & u_{1,2} & \cdots & u_{1,d} \\ 0 & u_{2,2} & \cdots & u_{2,d} \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & u_{d,d} \end{pmatrix} \quad L = \begin{pmatrix} l_{1,1} & 0 & \cdots & 0 \\ l_{2,1} & l_{2,2} & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ l_{d,1} & l_{d,2} & \cdots & l_{d,d} \end{pmatrix}$$

- ▶ Symmetric matrices : $A = A^T$.
- ▶ Unitary matrices : $AA^T = A^T A = I$ (matrix of an orthonormal basis).

Inverting a matrix

- ▶ A is diagonal, lower or upper triangular then :
 A invertible $\Leftrightarrow \prod_{i=1}^d a_{i,i} \neq 0$
- ▶ Lower triangular systems

$$Ax = b \Leftrightarrow \begin{pmatrix} a_{1,1} & 0 & \cdots & 0 \\ a_{2,1} & a_{2,2} & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ a_{d,1} & a_{d,2} & \cdots & a_{d,d} \end{pmatrix} \begin{pmatrix} x_1 \\ x_2 \\ \vdots \\ x_d \end{pmatrix} = \begin{pmatrix} b_1 \\ b_2 \\ \vdots \\ b_d \end{pmatrix} \text{ wh. } \prod_{i=1}^d a_{i,i} \neq 0$$

are solved recursively from the first to the last equation :

$$\left\{ \begin{array}{rcl} & a_{1,1}x_1 & = b_1 \\ & a_{2,2}x_2 + a_{2,1}x_1 & = b_2 \\ & a_{3,3}x_3 + a_{3,2}x_2 + a_{3,1}x_1 & = b_3 \\ & \vdots & \vdots \\ a_{1,1}x_1 + a_{1,2}x_2 + \cdots + a_{1,d}x_d & = & b_d \end{array} \right.$$

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$$\left\{ \begin{array}{rcl} & & x_1 = b_1/a_{1,1} \\ & & a_{2,1}x_1 + a_{2,2}x_2 = b_2 \\ & & a_{3,1}x_1 + a_{3,2}x_2 + a_{3,3}x_3 = b_3 \\ & & \vdots \\ a_{1,1}x_1 + a_{1,2}x_2 + \cdots + a_{1,d}x_d & = & b_d \end{array} \right.$$

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are solved recursively from the first to the last equation :

$$\left\{ \begin{array}{l} x_1 = b_1/a_{1,1} \\ x_2 = (b_2 - a_{2,1}b_1/a_{1,1})/a_{2,2} \\ a_{3,1}x_1 + a_{3,2}x_2 + a_{3,3}x_3 = b_3 \\ \vdots \\ a_{1,1}x_1 + a_{1,2}x_2 + \cdots + a_{1,d}x_d = b_d \end{array} \right.$$

Matrix determinant

▶ $A = \begin{pmatrix} a & b \\ c & d \end{pmatrix}$ is invertible iff $ad - bc \neq 0$ and $A^{-1} = \frac{1}{ad - bc} \begin{pmatrix} d & -b \\ -c & a \end{pmatrix}$

▶ For lower/upper triangular and diagonal matrices :

A is invertible iff $\prod_{i=1}^d a_{i,i} \neq 0$.

▶ In general, $A \in \mathbb{R}^{d \times d}$ is invertible

⇔ its d row (resp. column) vectors are linearly independent.

⇔ its determinant $\det(A) = \begin{vmatrix} a_{1,1} & a_{1,2} & \cdots & a_{1,d} \\ a_{2,1} & a_{2,2} & \cdots & a_{2,d} \\ \vdots & \vdots & \ddots & \vdots \\ a_{d,1} & a_{d,2} & \cdots & a_{d,d} \end{vmatrix} \neq 0$.

▶ The determinant is found recursively, developing on any row or column : $\det(A) = \sum_{i=1}^d a_{i,j} \text{Cof}(A)_{i,j}$.

▶ $\text{Cof}(A)_{i,j} = \det((a_{k,l})_{k \in \{1 \dots d\} \setminus \{i\}, l \in \{1 \dots d\} \setminus \{j\}})$

▶ if $\det(A) \neq 0$ then $A^{-1} = \frac{1}{\det(A)} \text{Cof}(A)^T$.

Eigenvalues, eigenvectors

A a square matrix.

Definition (Eigenvalues and eigenvectors)

- ▶ λ is an **eigenvalue** of A if there exists a vector $v \in \mathbb{R}^d$, $v \neq 0$ s.t. $Av = \lambda v$.
- ▶ Equivalently : λ is an **eigenvalue** of A if $\det(A - \lambda I) = 0$.
- ▶ Any v verifying $Av = \lambda v$ is an **eigenvector** associated to the eigenvalue λ .

- ▶ Properties :
 - ▶ For diagonal matrices, the eigenvalues are the diagonal elements (not for triangular matrices !).
 - ▶ 0 is an eigenvalue iff A is not invertible.
- ▶ A is **diagonalizable** if there exists a basis of eigenvectors :

$$A = PDP^{-1} \text{ with } D \text{ diagonal.}$$

Singular value decomposition

Symmetric matrices and eigenvalues/eigenvectors :

- ▶ A symmetric matrix is diagonalizable on an orthonormal basis :

$$A = PDP^T \text{ with } D \text{ diagonal, } PP^T = I.$$

- ▶ A symmetric matrix is said
 - ▶ **semi-definite positive** if $\langle x, Ax \rangle \geq 0, \forall x$.
Its eigenvalues are ≥ 0 .

*Any diagonal matrix with $\lambda_i \geq 0$,
 $A = B^T B$ for any $B \in \mathbb{R}^{m,d}$.*

- ▶ **definite positive** if $\langle x, Ax \rangle \geq 0, \forall x$ and $\langle x, Ax \rangle = 0, \Rightarrow x = 0$.
Its eigenvalues are > 0 .

*Any diagonal matrix with $\lambda_i > 0$,
 $A = B^T B$ for any $B \in \mathbb{R}^{m,d}$ when A is invertible.*

Note : a definite positive matrix defines a new norm on \mathbb{R}^d via the scalar product $\langle x, x \rangle_A = \langle x, Ax \rangle$

Singular value decomposition

Symmetric matrices and eigenvalues/eigenvectors :

- ▶ A symmetric matrix is diagonalizable on an orthonormal basis :

$$A = PDP^T \text{ with } D \text{ diagonal, } PP^T = I.$$

- ▶ A symmetric matrix is said
 - ▶ **semi-definite positive** if $\langle x, Ax \rangle \geq 0, \forall x$.
Its eigenvalues are ≥ 0 .

*Any diagonal matrix with $\lambda_i \geq 0$,
 $A = B^T B$ for any $B \in \mathbb{R}^{m,d}$.*

- ▶ **definite positive** if $\langle x, Ax \rangle \geq 0, \forall x$ and $\langle x, Ax \rangle = 0, \Rightarrow x = 0$.
Its eigenvalues are > 0 .

*Any diagonal matrix with $\lambda_i > 0$,
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Note : a definite positive matrix defines a new norm on \mathbb{R}^d via the scalar product $\langle x, x \rangle_A = \langle x, Ax \rangle$

Singular value decomposition

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Singular value decomposition

Fix $B \in \mathbb{R}^{m \times d}$, note that :

- ▶ $B^T B \in \mathbb{R}^{d \times d}$ and $BB^T \in \mathbb{R}^{m \times m}$ are symmetric semi-definite positive :
 - ▶ $B^T B = V \Delta_1 V^T$ with Δ_1 diagonal, $VV^T = I$ in $\mathbb{R}^{d \times d}$.
 - ▶ $BB^T = U \Delta_2 U^T$ with Δ_2 diagonal, $UU^T = I$ in $\mathbb{R}^{m \times m}$.
- ▶ One can show :
 - ▶ Δ_1 and Δ_2 have the same non-zero values $\lambda_1^2, \dots, \lambda_k^2$.
 - ▶ $B = UDV^T$ with

$$D = \text{diag}(\lambda_1, \dots, \lambda_k) = \begin{pmatrix} \lambda_1 & 0 & \dots & 0 & 0 & \dots & 0 \\ 0 & \lambda_2 & \dots & 0 & 0 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots & 0 & \dots & 0 \\ 0 & 0 & \dots & \lambda_k & 0 & \dots & 0 \\ 0 & 0 & \dots & 0 & 0 & \dots & 0 \end{pmatrix} \in \mathbb{R}^{m,d}.$$

- ▶ $B^T = VDU^T$ with $D = \text{diag}(\lambda_1, \dots, \lambda_k) \in \mathbb{R}^{d,m}$.
- ▶ $B = UDV^T$ is its singular value decomposition and $\lambda_1, \dots, \lambda_k$ its singular values.

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Other decompositions

▶ LU factorization

- ▶ for a diagonally dominant matrix A ($|a_{i,i}| \geq \sum_{j \neq i} |a_{i,j}|$)
- ▶ $A = LU$, L is lower triangular, U is upper triangular with 1 on the diagonal.
- ▶ $Ax = B$ solved in two steps : $Lz = b$ and $Ux = z$!

▶ Choleski decomposition

- ▶ for symmetric semi-definite positive matrices
- ▶ $A = U^T U$ with U upper triangular
- ▶ again easy to solve $Ax = b$ in two steps.

▶ QR decomposition

- ▶ for any matrix $A \in \mathbb{R}^{m \times d}$
- ▶ $A = QR$ with Q unitary in $\mathbb{R}^{m \times m}$ and R upper triangular.

Framework

▶ Random Space

- ▶ Ω is the set of random events.

$$\Omega = \{heads, tails\}$$

- ▶ \mathcal{A} is the set of “measurable” collections of events.

$$\mathcal{A} = \{\emptyset, \{heads\}, \{tails\}, \{heads, tails\}\}$$

- ▶ $\mathbb{P} : \mathcal{A} \rightarrow [0, 1]$ is the probability.

$$\begin{aligned} \mathbb{P}(\emptyset) &= 0, & \mathbb{P}(\{heads\}) &= p, \\ \mathbb{P}(\{tails\}) &= 1 - p, & \mathbb{P}(\{heads, tails\}) &= 1 \end{aligned}$$

▶ Properties of \mathbb{P}

- ▶ $0 \leq \mathbb{P} \leq 1$,
- ▶ $\mathbb{P}(\emptyset) = 0$, $\mathbb{P}(\Omega) = 1$,
- ▶ $A, B \in \mathcal{A}$, $A \cup B = \emptyset \Rightarrow \mathbb{P}(A \cup B) = \mathbb{P}(A) + \mathbb{P}(B)$ (chain rule).
- ▶ Equivalently : $A, B \in \mathcal{A}$, $\mathbb{P}(A \cup B) + \mathbb{P}(A \cap B) = \mathbb{P}(A) + \mathbb{P}(B)$.

- ▶ Random events are observed only through measurable quantities called **random variables**.

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Random variables

- ▶ A **Random variable** is a measurable function $X : (\Omega, \mathcal{A}) \rightarrow (\mathcal{F}, \mathcal{B}(\mathcal{F}))$
 - ↪ the measurability means $F \subset \mathcal{F} \Rightarrow X^{-1}(F) \subset \mathcal{A}$.
- ▶ $X(\Omega) \subset \mathcal{F}$ may be
 - ▶ finite ($\{0, 1\}$) or infinite (\mathbb{R}), discrete (\mathbb{N}) or continuous (\mathbb{R})
discrete/continuous random variables
 - ▶ have one or several variables (\mathbb{R}^d)
random variables/ random vectors.
- ▶ The measurability of X implies that \mathbb{P} may be transported to \mathcal{F} through X :

$$\mathbb{P}(\{\omega / X(\omega) \in F\}) = \mathbb{P}(X \in F) \stackrel{\text{def}}{=} \mathbb{P}_X(F)$$

\mathbb{P} is a probability on (Ω, \mathcal{A})
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Discrete random variables

Examples

- ▶ A single coin toss is a **Bernoulli variable** with parameter p
 - ▶ $X : (\Omega, \mathcal{A}) \rightarrow (\{0, 1\}, 2^{\{0,1\}})$,
 - ▶ $\mathbb{P}(X = 1) = p$, (hence $\mathbb{P}(X = 0) = 1 - p$).
 - ▶ Notation : $X \sim B(p)$.

- ▶ The sum of n independent coin tosses is a **multinomial** with parameter n, p
 - ▶ $Y : (\Omega, \mathcal{A}) \rightarrow (\{0, 1, \dots, n\}, 2^{\{0,1, \dots, n\}})$,
 - ▶ $Y = X_1 + X_2 + \dots + X_n$ where the X_i are independent copies $\equiv B(p)$.
 - ▶ $\mathbb{P}(Y = k) = \binom{n}{k} p^k (1 - p)^{n-k}$ for $k = 0 \dots n$.
 - ▶ Notation : $Y \sim Bin(n, p)$.

Discrete random variables

- ▶ \mathcal{F} is discrete $\mathcal{F} = \{x_1, x_2, \dots, x_N\}$, N finite or not.
- ▶ $X : (\Omega, \mathcal{A}) \rightarrow (\mathcal{F}, 2^{\mathcal{F}})$,
 - ▶ Notation : $\mathbb{P}(X = x_i) = p_i$. Note that $p_i \geq 0$ and $\sum_{i=1}^N p_i = 1$.
- ▶ The mean value or **expectation** of X is :

$$\begin{aligned}\mathbb{E}[X] &= \sum_{\omega \in \Omega} X(\omega) \mathbb{P}(\omega) \\ \mathbb{E}[X] &= \sum_{i=1}^N x_i \mathbb{P}_X(x_i)\end{aligned}$$

$$\text{Here, } \mathbb{E}[X] = \sum_{i=1}^N x_i p_i$$

- ▶ The **variance** of X is its deviation from its mean :

$$\begin{aligned}\text{Var}[X] &= \mathbb{E}[(X - \mathbb{E}[X])^2] \\ \text{Var}[X] &= \mathbb{E}[X^2] - \mathbb{E}[X]^2\end{aligned}$$

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Discrete random variables

- ▶ More generally for any measurable function $f : \mathcal{F} \rightarrow \mathbb{R}^d$, the expectation of $f(X)$ is :

$$\begin{aligned}\mathbb{E}[f(X)] &= \sum_{\omega \in \Omega} f(x) \mathbb{P}(X(\omega) = x) \\ \mathbb{E}[f(X)] &= \sum_{i=1}^N f(x_i) \mathbb{P}_X(x_i)\end{aligned}$$

$$\text{Here, } \mathbb{E}[f(X)] = \sum_{i=1}^N f(x_i) p_i$$

Bernoulli variables

- ▶ $X \sim B(p)$, hence
 $\mathcal{F} = \{0, 1\}$, $p_1 = p$, $p_0 = 1 - p$.

- ▶ The **expectation** of X is :

$$\begin{aligned}\mathbb{E}[X] &= \sum_{i=1}^N x_i p_i \\ \mathbb{E}[X] &= 0 * (1 - p) + 1 * p \\ \mathbb{E}[X] &= p\end{aligned}$$

- ▶ The **variance** of X is :

$$\begin{aligned}\text{Var}[X] &= \sum_{i=1}^N x_i^2 p_i - (\sum_{i=1}^N x_i p_i)^2 \\ \text{Var}[X] &= 0^2(1 - p) + 1^2 * p - p^2 \\ \text{Var}[X] &= p(1 - p).\end{aligned}$$

- ▶ The **expectation** of $f(X)$ is :

$$\begin{aligned}\mathbb{E}[f(X)] &= \sum_{i=1}^N f(x_i) p_i \\ \mathbb{E}[f(X)] &= f(0) * (1 - p) + f(1) * p.\end{aligned}$$

Discrete random vectors

- ▶ X has d coordinates, each of which is a discrete variable.

$$X = (X_1, \dots, X_d)^T : (\Omega, \mathcal{A}) \rightarrow (\mathcal{F} = \mathcal{F}_1 \times \dots \times \mathcal{F}_d, 2^{\mathcal{F}}),$$

- ▶ $\mathbb{P}(X = x_i) = p_i \Leftrightarrow \mathbb{P}(X = (x^1, \dots, x^d))$, where $x^i \in \mathcal{F}_i$.

- ▶ The **expectation** of X is the vector of the expectation of each coordinate :

$$\mathbb{E}[X] = (\mathbb{E}[X_1], \dots, \underset{\substack{\uparrow \\ \text{row } i}}{\mathbb{E}[X_i]}, \dots, \mathbb{E}[X_d])^T$$

- ▶ The variance is replaced by the **covariance matrix** :

- ▶ $\text{Cov}(X)$ is a $d \times d$ -matrix.
- ▶ $\text{Cov}(X)_{i,i} = \text{Var}(X_i)$.
- ▶ If $i \neq j$, $\text{Cov}(X)_{i,j} = \text{Cov}(X_i, X_j) = \mathbb{E}[X_i X_j] - \mathbb{E}[X_i] \mathbb{E}[X_j]$.

Discrete random vectors

Example

- ▶ $X = (X_1, X_2)$ with
 - ▶ $X_1 \sim B(p_1)$,
 - ▶ $X_2 \sim B(p_2)$,
 - ▶ X_1 and X_2 are decorrelated i.e. $\text{Cov}(X_1, X_2) = 0$.
- ▶ The **expectation** of X is :

$$\mathbb{E}[X] = \begin{pmatrix} \mathbb{E}[X_1] \\ \mathbb{E}[X_2] \end{pmatrix} = \begin{pmatrix} p_1 \\ p_2 \end{pmatrix}$$

- ▶ The **covariance matrix** of X is :

$$\text{Cov}[X] = \begin{pmatrix} \text{Var}[X_1] & \text{Cov}[X_1, X_2] \\ \text{Cov}[X_2, X_1] & \text{Var}[X_2] \end{pmatrix} = \begin{pmatrix} p_1(1-p_1) & 0 \\ 0 & p_2(1-p_2) \end{pmatrix}$$

Note : independence \Rightarrow decorrelation but the converse is false !

Continuous random variables

Real random variables

- ▶ $X : (\Omega, \mathcal{A}) \rightarrow (\mathbb{R}, \mathcal{B}(\mathbb{R}))$.
- ▶ $\mathbb{P}(X = x_i) = p_i \Leftrightarrow \mathbb{P}(X \in [a, b]) = P_X([a, b])$.
Note : $P_X \geq 0$ and $\int_{\mathbb{R}} dP_X(x) = 1$.
- ▶ The **expectations** and **variances** are defined as previously :

$$\begin{aligned}\mathbb{E}[X] &= \int_{\Omega} X(\omega) d\mathbb{P}(\omega) \\ \mathbb{E}[X] &= \int_{\mathbb{R}} x d\mathbb{P}_X(x)\end{aligned}$$

$$\begin{aligned}\mathbb{E}[f(X)] &= \int_{\Omega} f(X(\omega)) d\mathbb{P}(\omega) \\ \mathbb{E}[f(X)] &= \int_{\mathbb{R}} f(x) d\mathbb{P}_X(x)\end{aligned}$$

$$\mathbb{E}[\text{Var}(X)] = \mathbb{E}[X^2] - \mathbb{E}[X]^2$$

- ▶ If $d\mathbb{P}_X(x) = f_X(x)dx$ then f_X is the **probability density function of X (pdf)**.

Continuous random variables

Uniform distribution on $[a, b]$

- ▶ $X \sim \mathcal{U}_{[a,b]}$
- ▶ $\mathbb{E}[f(X)] = \int_{\mathbb{R}} f(x) dP_X(x) = \frac{1}{b-a} \int_{[a,b]} f(x) dx$
- ▶ pdf : $f_X(x) = \frac{1}{b-a} \delta_{[a,b]}(x)$

Gaussian distribution

of mean m and variance σ^2 :

- ▶ $X \sim \mathcal{N}_{m,\sigma^2}$
- ▶ $\mathbb{E}[f(X)] = \int_{\mathbb{R}} f(x) dP_X(x) = \int_{\mathbb{R}} f(x) \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{(x-m)^2}{2\sigma^2(x)}\right\} dx$
- ▶ pdf : $f_X(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{(x-m)^2}{2\sigma^2(x)}\right\}$

Continuous random variables

All we have seen previously extends to continuous random vectors such as :

Gaussian vector of mean \mathbf{m} and covariance matrix Σ^2 :

▶ $X = (X_1, \dots, X_d) \sim \mathcal{N}_{\mathbf{m}, \Sigma^2}$

▶ pdf : $f_X(x) = \frac{1}{(2\pi \det(\Sigma))^{d/2}} \exp \left\{ -\frac{(x-\mathbf{m})^T \Sigma^{-1} (x-\mathbf{m})}{2} \right\}$

$$\begin{aligned} \mathbb{E}[f(X)] &= \int_{\mathbb{R}^d} f(x_1, \dots, x_d) dP_X(x_1, \dots, x_d) \\ &= \int_{\mathbb{R}^d} f(x) * \frac{1}{(2\pi \det(\Sigma))^{d/2}} \exp \left\{ -\frac{(x-\mathbf{m})^T \Sigma^{-1} (x-\mathbf{m})}{2} \right\} dx \end{aligned}$$

Joint probabilities

Two simultaneous coin tosses :

- ▶ Each coin is fair $\mathbb{P}(\text{heads}) = \frac{1}{2}$
- ▶ All the possible outcomes of both draws ($\{\text{heads}, \text{heads}\}, \{\text{heads}, \text{tails}\}, \{\text{tails}, \text{heads}\}, \{\text{tails}, \text{tails}\}$) are equiprobable with $\mathbb{P}(\{\text{heads}, \text{heads}\}) = \frac{1}{4}$.
- ▶ Consider $Z = (X_1, X_2)$, X_i the random variable for tossing coin i . This means that :

$$\mathbb{P}(Z \in A \times B) = \mathbb{P}(X_1 \in A)\mathbb{P}(X_2 \in B)$$

or in other words :

$$P_{(X_1, X_2)} = P_{X_1} P_{X_2}$$

X_1 and X_2 are **independent**.

Joint probabilities

But this is not always the case :

Example

X/Y	Sick	Fit	Total
Positive test	90	100	190
Negative test	10	900	910
Total	100	1000	1100

- ▶ $\mathbb{P}(X = \textit{positive}) = 190/1100$
- ▶ $\mathbb{P}(Y = \textit{sick}) = 100/1100$
- ▶ Clearly :

$$\mathbb{P}((X, Y) = (\textit{positive}, \textit{sick})) = 90/1100$$

\neq

$$\mathbb{P}(X = \textit{positive})\mathbb{P}(Y = \textit{sick}) = 100 * 190/1100^2$$

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Independence

Definition (Independence)

X and Y are independent random variables ($X \perp\!\!\!\perp Y$) if and only if their joint probability $\mathbb{P}_{X,Y}$ is the product of their marginal probabilities : $\mathbb{P}_{X,Y} = \mathbb{P}_X \mathbb{P}_Y$.

Also, X_1, \dots, X_n are independent iff $\mathbb{P}_{X_1, \dots, X_n} = \prod_{i=1}^n P_{X_i}$.

► Equivalently :

- $\forall A, B \quad \mathbb{P}((X, Y) \in A \times B) = \mathbb{P}(X \in A) \mathbb{P}(Y \in B)$
- $\forall f, g \quad \mathbb{E}[f(X)g(Y)] = \mathbb{E}[f(X)] \mathbb{E}[g(Y)]$

► If X and Y are independent then $\text{Cov}(X, Y) = 0$.

► For Gaussian variables only : $\text{Cov}(X, Y) = 0 \Leftrightarrow X \perp\!\!\!\perp Y$.

If X and Y are independent, knowing X does not give any information on Y , what if they are not independent ?

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Conditional probabilities

Example

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Positive test	90	100	190
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- ▶ Amongst all people :

$$\mathbb{P}(Y = \textit{sick}) = 100/1100,$$

$$\mathbb{P}(Y = \textit{fit}) = 1000/1100$$

- ▶ Amongst people with a positive test :

$$\mathbb{P}(Y = \textit{sick} | X = \textit{positive}) = 90/190,$$

$$\mathbb{P}(Y = \textit{fit} | X = \textit{positive}) = 100/190,$$

- ▶ Amongst people with a negative test :

$$\mathbb{P}(Y = \textit{sick} | X = \textit{negative}) = 10/910,$$

$$\mathbb{P}(Y = \textit{fit} | X = \textit{negative}) = 900/910,$$

Conditional probabilities

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Negative test	10	900	910
Total	100	1000	1100

- ▶ Amongst all people :

$$\mathbb{P}(Y = \textit{sick}) = 100/1100,$$

$$\mathbb{P}(Y = \textit{fit}) = 1000/1100$$

- ▶ Amongst people with a positive test :

$$\mathbb{P}(Y = \textit{sick} | X = \textit{positive}) = 90/190,$$

$$\mathbb{P}(Y = \textit{fit} | X = \textit{positive}) = 100/190,$$

- ▶ Amongst people with a negative test :

$$\mathbb{P}(Y = \textit{sick} | X = \textit{negative}) = 10/910,$$

$$\mathbb{P}(Y = \textit{fit} | X = \textit{negative}) = 900/910,$$

Conditional probabilities

Example

X/Y	Sick	Fit	Total
Positive test	90	100	190
Negative test	10	900	910
Total	100	1000	1100

- ▶ Amongst people with a positive test :

$$\mathbb{P}(Y = \textit{sick} | X = \textit{positive}) = 90/190,$$

$$\mathbb{P}(Y = \textit{fit} | X = \textit{positive}) = 100/190,$$

- ▶ Note :

$$\mathbb{P}(Y = \textit{sick} | X = \textit{negative})\mathbb{P}(X = \textit{negative}) = \mathbb{P}((Y, X) = (\textit{sick}, \textit{negative})),$$

Definition (Conditional probabilities)

$$\mathbb{P}(A \textit{ and } B) = \mathbb{P}(A|B)\mathbb{P}(B) = \mathbb{P}(B|A)\mathbb{P}(A)$$

Conditional probabilities

More generally :

Definition

The conditional probability $\mathbb{P}_{X|Y}$ is the probability s.t. :

$$\forall f, \mathbb{E}[f(X, Y)] = \int f(X, Y) dP_{X, Y} = \int dP_Y \int f(X, Y) dP_{X|Y}$$

- ▶ For discrete random variables :

$$\mathbb{P}((X, Y) = (x, y)) = \mathbb{P}(Y = y | X = x) \mathbb{P}(X = x)$$

- ▶ If (X, Y) and Y have pdf $p_{(X, Y)}$ and p_Y , then $P_{X|Y}$ is a the corresponding pdf : $p_{X|Y} = \frac{p_{(X, Y)}}{p_Y}$
- ▶ $\mathbb{E}[X|Y]$ is the conditional esperance of X given Y is a random variable. It is the projection of X on the set of random variables of the form $g(Y)$.

Bayes rule, maximum likelihood, maximum a posteriori

Framework :

- ▶ Y is a random variable, Y is observed.
- ▶ Θ is a random variable, Θ is the parameter.
- ▶ Goal : given observed data Y , find the best guess for Θ .

Probabilities

- ▶ The conditional probability of the observations : $\mathbb{P}_{Y|\Theta}$.
- ▶ The prior : \mathbb{P}_{Θ} .
- ▶ The posterior : $\mathbb{P}_{\Theta|Y}$.

Bayes rule

$$\mathbb{P}_{\Theta|Y}(\theta, y) = \frac{\mathbb{P}_{Y|\Theta}(y, \theta) \mathbb{P}_{\Theta}(\theta)}{\int \mathbb{P}_{Y|\Theta}(y, \theta') \mathbb{P}_{\Theta}(\theta') d\theta}$$

Estimator

- ▶ Maximum likelihood : $\theta_{ML} = \operatorname{argmax}_{\theta} \mathbb{P}_{Y|\Theta}(y, \theta)$.
- ▶ Maximum a posteriori : $\theta_{MAP} = \operatorname{argmax}_{\theta} \mathbb{P}_{\Theta|Y}(\theta, y)$.
- ▶ Bayes mean square estimator : $\theta_M = \mathbb{E}[\Theta|Y]$.

Information theory

- ▶ **Entropy** measures the amount of disorder of X :
 - ▶ $H(X) = - \int P_X(x) \log(P_X(x)) dx$. Note : $H(X) \geq 0$.
 - ▶ For discrete random variables :
 - ▶ $X \sim \mathcal{U}$ maximizes the entropy $H = \log(N)$.
 - ▶ $X \sim \delta_{x_i}$ minimizes the entropy $H = \frac{1}{N} \log(N)$.
- ▶ The **Kullback-Leibler divergence** compares the laws of X and Y :
 - ▶ $D(X||Y) = \int P_X(x) \log \left(\frac{P_X(x)}{P_Y(x)} \right) dx$. Note : $D(X||Y) \neq D(Y||X)$.
 - ▶ $D(X||Y) \geq 0$ and $[D(X||Y) = 0 \Leftrightarrow P_X = P_Y]$.
- ▶ The **mutual information** measures the amount of shared information between X and Y :
 - ▶ $I(X, Y) = D(P_{(X, Y)} || P_X P_Y)$. Note : $I(X, Y) = I(Y, X)$.
 - ▶ $I(X, Y) \geq 0$ and $[I(X, Y) = 0 \Leftrightarrow X \perp\!\!\!\perp Y]$.
- ▶ The **perplexity** is a measure of complexity of a distribution :
 - ▶ $P(X) = 2^{H(X)}$.
 - ▶ this is a common way of evaluating language models.

Approximations and confidence intervals

- ▶ Statistical learning (classification) :
 - ▶ Goal : from i.i.d 1 samples $(x_i, y_i)_{i=1\dots n}$, find a hypothesis f that minimizes the risk : $\mathbb{E}[\text{loss}(f(X), Y)]$
 - ▶ $\mathbb{E}[\text{loss}(f(x), Y)]$ is not known, only its empirical version is accessible : $\frac{1}{n} \sum \text{loss}(f(x_i), y_i)$

↔ need to control how far is the empirical loss to the true one.

- ▶ Some tools to do so are :

- ▶ Markov inequality : $\mathbb{P}(X > \epsilon) \leq \frac{\mathbb{E}[X]}{\epsilon}$

- ▶ Chebicheff inequality : $\mathbb{P}(|X - \mathbb{E}[X]| \geq \epsilon) \leq \frac{\text{Var}[X]}{\epsilon^2}$

Apply this to $S_n = \frac{1}{n} \sum_{i=1}^n X_i$, with X_i i.i.d X , one gets :

$$\mathbb{P}(|S_n - \mathbb{E}[X]| \geq \epsilon) \leq \frac{\text{Var}[X]}{n\epsilon^2}$$

(S_n is the empirical risk, $\mathbb{E}[X]$ the true one.)

- ▶ Chernoff-Hoeffding bound : $\mathbb{P}(|S_n - \mathbb{E}[X]| \geq \epsilon) \leq e^{-2n\epsilon^2}$

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► Proof of Markov inequality

$$\mathbb{E}[X] = \int x d\mathbb{P}_X(x) = \int_{x \geq \epsilon} x d\mathbb{P}_X(x) + \int_{x < \epsilon} x d\mathbb{P}_X(x)$$

$$\mathbb{E}[X] \geq \int_{x \geq \epsilon} x d\mathbb{P}_X(x) \geq \epsilon \int_{x \geq \epsilon} d\mathbb{P}_X(x)$$

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► From bounds to confidence intervals

Chebicheff inequality : $\mathbb{P}(|S_n - \mathbb{E}[X]| \geq \epsilon) \leq \frac{\text{Var}[X]}{n\epsilon^2}$

► $\frac{\text{Var}[X]}{n\epsilon^2} \leq \delta$ implies : $\mathbb{P}(|S_n - \mathbb{E}[X]| \geq \epsilon) \leq \delta$ or

If $n \geq \frac{\text{Var}[X]}{\delta\epsilon^2}$ then with probability at least $1 - \delta$, $|S_n - \mathbb{E}[X]| \leq \epsilon$.

► Then if $n = \frac{\text{Var}[X]}{\delta\epsilon^2}$, we obtain :

For all n , with probability at least $1 - \delta$, $|S_n - \mathbb{E}[X]| \leq \sqrt{\frac{\text{Var}[X]}{n\delta}}$.

$$\mathbb{E}[\text{loss}(f(X), Y)] \in \mathbb{E}_{\text{emp}}[\text{loss}(f(X), Y)] + \left[-\sqrt{\frac{\text{Var}[X]}{n\delta}}, \sqrt{\frac{\text{Var}[X]}{n\delta}} \right]$$

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Minimizing a function

Goal : find the global minimum/minimizer of $f : \mathbb{R}^d \rightarrow \mathbb{R}$.

Potential problems / partial solutions :

- ▶ Existence of a global minimum ?
 - ↪ f is continuous and coercive ($f(x) \rightarrow \infty$ when $\|x\| \rightarrow \infty$).
- ▶ Characterization of the minimizers ?
 - ↪ f is C^1 . If x^* is a local minimizer then its gradient $\nabla f(x) = 0_{\mathbb{R}^d}$.
 - ↪ f is C^2 . x^* is a local minimizer iff its gradient $\nabla f(x) = 0_{\mathbb{R}^d}$ and its hessian $\nabla^2 f(x)$ is a semi-definite positive matrix.
- ▶ Characterization of the global minimizers ?

Zeroing the gradient is not sufficient (maxima, saddle points,...) !

Minimizing a function

Goal : find the global minimum/minimizer of $f : \mathbb{R}^d \rightarrow \mathbb{R}$ for $x \in Q$.

- ▶ Constrained minimization ($Q \neq \mathbb{R}^d$) : characterization of the minimizers ?
 - ↪ minimizers may be on the border of Q : $\nabla f(x^*) \neq 0$!
- ▶ Gradient descents :
 - ▶ Algorithms of the form : $x^{t+1} = x^t - \gamma_t \nabla f(x^t)$
 - ▶ Ex : Gauss-Newton, conjugate gradient descent,...
 - ▶ Convergence ?
- ▶ What if f is not differentiable ?

Convex functions

Definition (convex functions)

$f : \mathbb{R}^d \rightarrow \mathbb{R}$ is convex iff $\forall \lambda \in [0, 1], \forall x, y \in \mathbb{R}^d,$
$$f(\lambda x + (1 - \lambda)y) \leq \lambda f(x) + (1 - \lambda)f(y)$$

$f : \mathbb{R}^d \rightarrow \mathbb{R}$ is strictly convex iff $\forall \lambda \in [0, 1], \forall x, y \in \mathbb{R}^d, \text{ s.t } x \neq y$
$$f(\lambda x + (1 - \lambda)y) < \lambda f(x) + (1 - \lambda)f(y)$$

- ▶ Other characterizations
 - ▶ If $f \in C^2$, f convex iff its $\nabla^2 f$ is semi-definite positive.
 - ▶ $f : \mathbb{R}^d \rightarrow \mathbb{R}$, f convex iff f' is non-decreasing iff $f'' \geq 0$.
 - ▶ f lies over all its tangents.
- ▶ *Ex. : affine functions, square loss, exp,...*
- ▶ Properties
 - ▶ no maxima, no saddle points and non local minima !
 - ▶ $\nabla f(x) = 0 \Rightarrow x$ is a global minimizer.

Convex functions are easier to minimize !

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A convex and constrained problem in classification

Problem

- ▶ Inputs : $\{x_i, y_i\}_{i=1..n}$, $x_i \in \mathbb{R}^d$, $y_i \in \{0, 1\}$.
- ▶ Goal : (P) Min $J(w, b) = \frac{1}{2}\|w\|^2 + \sum_1^n \max(0, 1 - y_i(wx_i + b))$

Resolution :

- ▶ Rewrite (P) as :
Min $J(w, b, \xi) = \frac{1}{2}w^2 + \sum_1^n \xi_i$ s.t. $y_i(wx_i + b) \geq 1 - \xi_i$ and $\xi_i \geq 0$
- ▶ Introduce a Lagrange multiplier for each constraint :
 $L(w, b, \xi, \alpha, \eta) = \frac{1}{2}\|w\|^2 + \sum_1^n \xi_i + \sum_i \alpha_i(1 - \xi_i - y_i(wx_i + b)) + \sum_i \eta_i \xi_i$,
 $\alpha_i \geq 0, \eta_i > 0$.
- ▶ The first order conditions $\partial_w J = 0, \partial_\xi J = 0, \partial_b J = 0$ yield :
 $w = \sum_i \alpha_i y_i x_i \quad \sum_i \alpha_i y_i = 0 \quad \forall i, 1 = \alpha_i + \eta_i$
- ▶ Which substituted in (P) gives the dual problem :
Maximize $J(\alpha) = \frac{1}{2}\|\sum_i \alpha_i y_i x_i\|^2 - \alpha^T \mathbf{1}$ s.t. $0 \leq \alpha \leq 1$